Machine Learning Methods in Empirical Finance

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Lecture 1 XVIII Encontro Brasileiro de Finanças

Introduction

► What is Machine Learning?

What is Machine Learning? – What do we want to learn?

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- What do we want to learn?
- From what do we want to
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- ► When ML methods are statistically sound they are called Statistical Learning (SL) methods.

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- Causal inference is a goal for decision making.

A great matching: Machine learning with Big Data A great matching: Machine learning with Big Data with Econometrics

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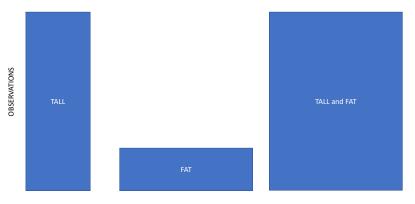
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- A quote from SAS (www.sas.comen.us/insightsbig-datawhat-is-big-data.html): "Big data is a term that describes the large volume of data – both structured and unstructured – that inundates a business on a day-to-day basis. But it's not the amount of data that's important. It's what organizations do with the data that matters. Big data can be analyzed for insights that lead to better decisions and strategic business moves."



VARIABLES



Source: https://solutionsreview.com

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- ▶ Unstructured data has internal structure but is not organized via pre-defined data models or schema.
- ▶ Examples: text files, web pages, social media, email, etc...

From unstructured to structured data

Example: Economic Policy Uncertainty

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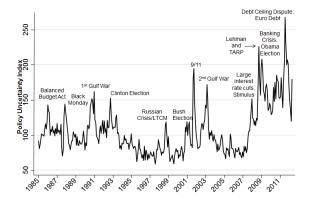
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- ► From unstructured to structured data: The first component is an index of search results from 10 large newspapers. Normalized index of the volume of news articles discussing economic policy uncertainty.

USA Today, the Miami Herald, the Chicago Tribune, the Washington Post, the Los Angeles Times, the Boston Globe, the San Francisco Chronicle, the Dallas Morning News, the New York Times, and the Wall Street Journal.

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Example: Economic Policy Uncertainty



Source: http://www.policyuncertainty.com and Baker, Bloom and Davis(QJE, 2016).

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Example: News implied VIX (NVIX) Moreira and Manela (JFE, 2017)

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- ► In US postwar data, periods when NVIX is high are followed by periods of above average stock returns, even after controlling for contemporaneous and forward-looking measures of stock market volatility.
- ▶ NVIX is a key predictor of the equity premium.
- ▶ Methodology: ML regression of VIX on regressors based on text data.

From unstructured to structured data

Example: News implied VIX (NVIX)

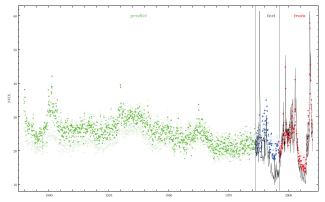


Fig. 1. News implied volatility 1880–2009. Solid line is end-of-month Chicago Board Options Exchange volatility implied by potions XVD, Dots are news implied volatility (NVX, VVX, v = v = v = v = v = v = parances of - para i in month r scaled by total month r n-grams and w is estimatedwith a support vector regression. The mini subsample, 1996 to 2009, is used to estimate the dependency between news data and implied volatility. Theter subsample, 1996–1995, is used for our-of-sample tests of model fit. The parefer subsample includes all earlier dowerations for which options data $and, hence, VX are not available. Light-colored triangles indicate a nonparametric bootstrap 953 confidence interval around <math>V\overline{X}$ using one thousand randomizations. These show the essistivity of the predicted values to randomizations of the rain subsample.

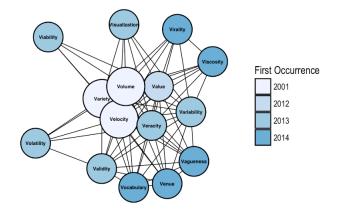
Source: Moreira and Manela (JFE, 2017).

What is "Big Data"? The Vs of "Big Data"



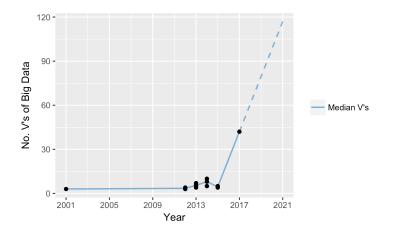
Source: http://www.ibmbigdatahub.com

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- ▶ Lots of potential applications due to availability of massive datasets and new tools.

Models/Methods

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- ► Linear regression is a **GREAT** ML method!

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What is a Machine Learning Model?

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- The text is decomposed into five categories: War, Financial Intermediation, Government, Stock Markets, and Natural Disasters.

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 - Models: linear regression, additive models, regression trees, random forests, neural networks, deep learning, kernel regression, series regression, splines.

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► Unsupervised learning:

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- The goal is to find "interesting" patters in data and there are no desired outputs given a set of inputs.

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► Unsupervised learning:

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- The goal is to find "interesting" patters in data and there are no desired outputs given a set of inputs.
- Unconditional models, cluster analysis, missing value imputation, factor construction, etc.

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Big Data + Big Models + Big Set of Models

BIG PROBLEM!!!!

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- ▶ The lack of uniform convergence is not a problem of Big Data (high-dimensions) and it is due to the model search methods that are applied before inference is conducted.
- ▶ On the other hand, prediction (forecasting) after model selection is a much easier task.

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- **High-Dimension**: Models with more candidate variables than observations, and the number of candidate variables grows polynomially or exponentially with n (or T).

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 - 3. Bayesian methods.

The Road Map

Lecture 1:

- ► Linear models with shrinkage
- ▶ Applications to covariance matrix forecasting

Lecture 2:

- ▶ Nonlinear models
- ► Applications to equity premium forecasting

Shrinkage in Linear Models: Ridge, LASSO, Adaptive LASSO, Elastic Net

What happens when p >> T in linear regressions?

▶ We are interested in single-equation linear models

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- $\boldsymbol{x}_t = [\boldsymbol{x}_t(S)', \boldsymbol{x}_t(S^c)']', \boldsymbol{x}_t(S) \in \mathbb{R}^s$ is the vector of relevant variables and $\boldsymbol{x}_t(S^c) \in \mathbb{R}^{p-s}$ is the vector of irrelevant ones. $\boldsymbol{\beta} = [\boldsymbol{\beta}'_S, \boldsymbol{\beta}'_{S^c}]'.$

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► Goals:

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- 2. Estimate β_S as if the correct set of variables is known to the econometrician.

• A Penalized Least Squares estimator $\hat{\beta}$:

$$\left| \widehat{oldsymbol{eta}}(\lambda) = rgmin_{oldsymbol{eta}\in\mathcal{B}} \sum_{t=1}^T (y_t - oldsymbol{eta}' oldsymbol{z}_t)^2 + \sum_{j=1}^p p_{\lambda}(|oldsymbol{eta}_j|),
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- ► Key assumption (for some methods): **sparsity**.

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- Sparse modeling has been successfully used to deal with high-dimensional models and is a crucial condition for identifiability.

$$\widehat{\boldsymbol{\beta}}_{Ridge}(\lambda) = \operatorname*{arg\,min}_{\boldsymbol{\beta}\in\mathcal{B}} \sum_{t=1}^{T} (y_t - \boldsymbol{\beta}' \boldsymbol{z}_t)^2 + \lambda \sum_{j=0}^{p} \beta_j^2$$

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▶ Good for prediction but not for variable selection.

$$\widehat{\boldsymbol{\beta}}_{LASSO}(\lambda) = \operatorname*{arg\,min}_{\boldsymbol{\beta} \in \mathcal{B}} \sum_{t=1}^{T} (y_t - \boldsymbol{\beta}' \boldsymbol{z}_t)^2 + \lambda \sum_{j=1}^{p} |\beta_j|$$

Least Absolute Shrinkage and Selection Operator (LASSO):

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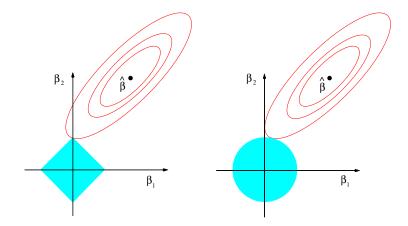
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- ▶ The regularization path can be efficiently estimated.
- Can handle (many) more variables than observations (p >> T).
- ▶ Under some conditions can select the correct subset of relevant variables.

LASSO versus Ridge



LASSO and Model Selection

Consistency

Estimation Consistency

$$\widehat{\boldsymbol{\beta}} - \boldsymbol{\beta}^0 \stackrel{p}{\longrightarrow} \mathbf{0}, \text{ as } T \longrightarrow \infty.$$

Model Selection Consistency

$$\mathsf{P}\left(\left\{i:\widehat{\boldsymbol{\beta}}\neq\mathbf{0}\right\}=\left\{i:\boldsymbol{\beta}^{0}\neq\mathbf{0}\right\}\right)\longrightarrow1,\,\mathrm{as}\,T\longrightarrow\infty.$$

Sign Consistency

$$\mathsf{P}\left(\widehat{\boldsymbol{\beta}} \stackrel{s}{=} \boldsymbol{\beta}^{0}\right) \longrightarrow 1 \text{ as } T \longrightarrow \infty$$

where

$$\widehat{oldsymbol{eta}} \stackrel{s}{=} oldsymbol{eta}^0 \Longleftrightarrow \mathsf{sign}\left(\widehat{oldsymbol{eta}}
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LASSO and Model Selection $$\operatorname{The\ sign\ Function}$$

The sign function is defined as

$$\mathsf{sign}(x) = \begin{cases} 1 & \text{ if } x > 0 \\ 0 & \text{ if } x = 0 \\ -1 & \text{ if } x < 0 \end{cases}$$

Sign Consistency Definitions

Strong Sign Consistency The LASSO estimator is **strongly sign consistent** if $\exists \lambda_T = f(T)$ such that

$$\lim_{T \to \infty} \mathsf{P}\left(\widehat{\boldsymbol{\beta}}(\lambda_T) \stackrel{s}{=} \boldsymbol{\beta}^0\right) = 1$$

General Sign Consistency

The LASSO estimator is general sign consistent if

$$\lim_{T \to \infty} \mathsf{P}\left(\exists \lambda, \widehat{\boldsymbol{\beta}}(\lambda) \stackrel{s}{=} \boldsymbol{\beta}^0\right) = 1$$

► Strong sign consistency **implies** general sign consistency

LASSO and Model Selection Sign Consistency

General Sign Consistency versus Strong Sign Consistency

 Strong Sign Consistency implies one can use a pre-selected λ to achieve consistent model selection via the LASSO.

LASSO and Model Selection Sign Consistency

General Sign Consistency versus Strong Sign Consistency

- Strong Sign Consistency implies one can use a pre-selected λ to achieve consistent model selection via the LASSO.
- General Sign Consistency means for a random realization there exists a correct amount of regularization that selects the true model.

LASSO and Model Selection

Irrepresentable Condition

Strong Irrepresentable Condition $\exists \eta > 0 \text{ such that}$

$$\left| \widehat{\boldsymbol{\Sigma}}_{S^c S} \widehat{\boldsymbol{\Sigma}}_{SS}^{-1} \mathsf{sign} \left(\boldsymbol{\beta}_S^0 \right) \right| \leq 1 - \eta$$

Weak Irrepresentable Condition

$$\left| \widehat{\mathbf{\Sigma}}_{S^c S} \widehat{\mathbf{\Sigma}}_{SS}^{-1} \mathsf{sign} \left(oldsymbol{eta}_S^0
ight)
ight| < \mathbf{1}$$

- $\mathbf{1} \in \mathbb{R}^{(p-s)}$ is a vector of ones, and the inequality holds element-wise.
- ► The Irrepresentable Condition is a key condition for model selection consistency of the LASSO!

LASSO and Model Selection

Irrepresentable Condition

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- $\mathbf{1} \in \mathbb{R}^{(p-s)}$ is a vector of ones, and the inequality holds element-wise.
- ► The Irrepresentable Condition is a key condition for model selection consistency of the LASSO!
- ► This is a too strong condition!

 $\blacktriangleright\,$ The Adaptive LASSO (adaLASSO) estimator is given by

$$\widehat{\boldsymbol{\beta}}_{adaLASSO} = \operatorname*{arg\,min}_{\boldsymbol{\beta}\in\mathcal{B}} \sum_{t=1}^{T} (y_t - \boldsymbol{\beta}' \boldsymbol{z}_t)^2 + \lambda \sum_{j=1}^{p} \boldsymbol{w}_j |\beta_j|,$$

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where w_1, \ldots, w_p are non-negative pre-defined weights.

• Usually $w_j = |\tilde{\beta}_j|^{-\tau}$, for $\tau > 0$, where $\tilde{\beta}_j$ is an initial estimator (e.g., LASSO).

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- ► Usually $w_j = |\tilde{\beta}_j|^{-\tau}$, for $\tau > 0$, where $\tilde{\beta}_j$ is an initial estimator (e.g., LASSO).
- ▶ Provide consistent estimates for the non-zero parameters;
- ▶ Has the oracle property under some conditions.
- ► Theoretical results in general time-series framework: Medeiros and Mendes (JoE, 2016)

▶ The Naïve Elastic Net estimator is defined as

$$\widehat{\boldsymbol{\beta}}(\text{na\"ive}) = \operatorname*{arg\,min}_{\boldsymbol{\beta} \in \in \mathbb{R}^p} \sum_{t=1}^T (y_t - \boldsymbol{\beta}' \boldsymbol{z}_t)^2 + \lambda_2 \sum_{j=1}^p \beta_i^2 + \lambda_1 \sum_{j=1}^p |\beta_i|.$$

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► The Elastic Net estimator is given by

$$\widehat{\boldsymbol{\beta}} = (1 + \lambda_2)\widehat{\boldsymbol{\beta}}(\text{na\"ive}).$$

▶ The Naïve Elastic Net estimator is defined as

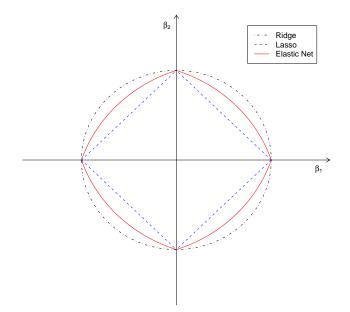
$$\widehat{\boldsymbol{\beta}}(\text{na\"ive}) = \operatorname*{arg\,min}_{\boldsymbol{\beta} \in \in \mathbb{R}^p} \sum_{t=1}^T (y_t - \boldsymbol{\beta}' \boldsymbol{z}_t)^2 + \lambda_2 \sum_{j=1}^p \beta_i^2 + \lambda_1 \sum_{j=1}^p |\beta_i|.$$

▶ The Elastic Net estimator is given by

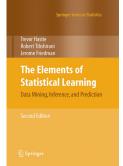
$$\widehat{\boldsymbol{\beta}} = (1 + \lambda_2)\widehat{\boldsymbol{\beta}}(\text{na\"ive}).$$

▶ The naïve EL-Net estimator selects the same model as the EL-Net version.

The Geometry of the Elastic Net



To Learn More about Shrinkage



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Empirical Example: Forecasting Large Dimensional Realized Covariance Matrices

Callot, Laurent, Anders B. Kock and Marcelo C. Medeiros (2017). *Modeling and Forecasting Large Realized Covariance Matrices and Portfolio Choice*. Journal of Applied Econometrics, 32, 140-158.

Dataset

- ► 30 stocks from the Dow Jones index from 2006 to 2012 with a total of 1474 daily observations.
- ▶ Daily realized covariances are constructed from 5 minutes returns by the method of Lunde, Shephard, Sheppard (2013).

Basic Materials 2	Technology 4	Consumer Cyclical 3	Consumer Non-cyclical 7
Energy	Financial	Industrial	Communication 4
2	3	5	

▶ The stocks can be classified in 8 broad categories.

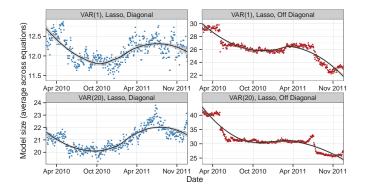
Results: Sectors

		Variance Equations							
	Basic Materials	0.75	0.40	0.14	0.52	0.23	0.35	0.57	0.39
Lagged covariance Lagged variance	Consumer, Non-cyclical	0.17	0.48	0.37	0.37	0.24	0.20	0.26	0.32
	Financial	0.00	0.42	0.99	0.24	0.64	0.20	0.12	0.48
	Communications	0.32	0.23	0.10	0.57	0.19	0.14	0.27	0.19
	Industrial	0.00	0.19	0.28	0.16	1.00	0.08	0.07	0.18
	Energy	0.58	0.45	0.46	0.33	0.02	1.00	0.38	0.55
	Technology	0.34	0.19	0.09	0.24	0.02	0.05	0.63	0.12
	Consumer, Cyclical	0.34	0.54	0.35	0.29	0.30	0.20	0.31	0.70
	Basic Materials	0.00	0.00	0.00	0.00	0.02	0.00	0.01	0.01
	Consumer, Non-cyclical	0.00	0.01	0.00	0.00	0.00	0.00	0.00	0.00
	Financial	0.00	0.01	0.02	0.00	0.00	0.00	0.01	0.00
	Communications	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
	Industrial	0.02	0.01	0.01	0.02	0.03	0.00	0.03	0.02
	Energy	0.01	0.03	0.01	0.03	0.01	0.01	0.02	0.03
	Technology	0.00	0.00	0.00	0.01	0.00	0.00	0.01	0.00
Т	Consumer, Cyclical	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00

Results: Sectors

		Covariance Equations							
	Basic Materials	0.81	0.27	0.10	0.24	0.24	0.93	0.26	0.29
ce	Consumer, Non-cyclical	0.48	0.71	0.56	0.32	0.28	0.35	0.36	0.41
ian	Financial	0.13	0.25	0.64	0.16	0.06	0.32	0.15	0.17
Lagged variance	Communications	0.62	0.57	0.54	0.65	0.51	0.70	0.61	0.58
	Industrial	0.13	0.13	0.18	0.18	0.34	0.08	0.10	0.07
	Energy	0.12	0.08	0.21	0.06	0.03	0.56	0.11	0.12
La	Technology	0.74	0.49	0.51	0.52	0.43	0.34	0.82	0.53
	Consumer, Cyclical	0.14	0.52	0.55	0.37	0.37	0.51	0.29	0.90
	Basic Materials	0.09	0.12	0.12	0.11	0.09	0.12	0.09	0.11
nce	Consumer, Non-cyclical	0.04	0.05	0.05	0.04	0.03	0.05	0.04	0.03
l covariance	Financial	0.11	0.14	0.18	0.11	0.07	0.10	0.10	0.10
	Communications	0.07	0.09	0.09	0.11	0.04	0.10	0.08	0.08
	Industrial	0.17	0.14	0.15	0.14	0.24	0.11	0.14	0.15
ge	Energy	0.16	0.16	0.16	0.15	0.10	0.24	0.15	0.16
Lagged	Technology	0.08	0.07	0.09	0.08	0.05	0.07	0.08	0.06
	Consumer, Cyclical	0.05	0.05	0.05	0.06	0.04	0.08	0.05	0.06

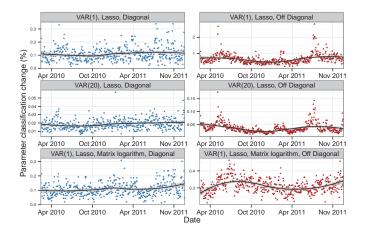
Results: Average Equation Size



- ▶ Diagonal equations more stable than off-diagonal ones.
- ▶ Diagonal equations are smaller.
- ► Flash Crash: May 6th 2010.

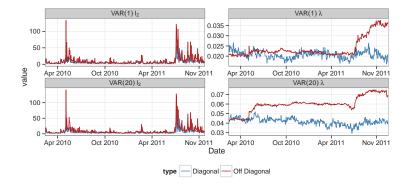
Results: Parameter Stability

Fraction of parameters that change from being zero to non-zero or vice versa in two consecutive periods

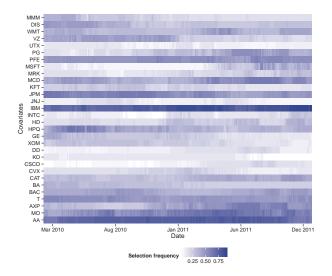


Results: Forecast Error and Penalty Parameter

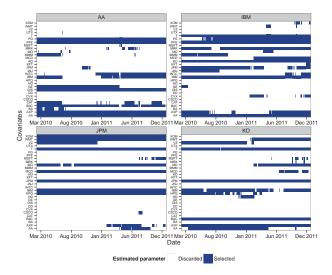
 ℓ_2 -norm of the 1-step ahead forecast error (left panel) and average penalty parameter (right panel) selected by BIC.



Results: Selection Frequency - VAR(1) LASSO

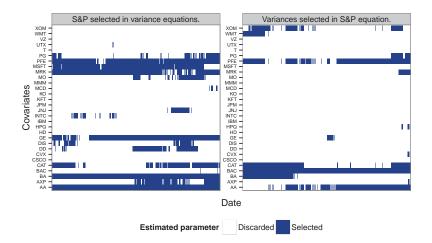


Results: Selection - VAR(1) LASSO



Results: Common Factor – VAR (1) LASSO

Lagged variance of the S&P selected in the variance equations of the Dow Jones stocks (left panel) and lagged variances of the Dow Jones stocks selected in the equation of the variance of the S&P 500 (right panel).



Forecasting Results

		AMedAFE		AMaxAFE			ℓ_2			
Model	h	Α	D	0	Α	D	0	Α	D	0
No-Change	1	0.33	0.57	0.33	3.53	3.53	1.47	11.22	5.98	9.22
Censored	5	0.46	0.79	0.45	4.51	4.51	1.91	15.02	7.89	12.41
	20	0.58	0.98	0.57	5.12	5.12	2.22	18.05	9.25	15.17
DCC	1	0.56	0.95	0.55	8.40	8.36	4.28	22.37	12.40	18.17
$\mathrm{EWMA}(\lambda=0.96)$	1	0.88	1.08	0.88	8.07	8.03	4.55	28.89	12.55	25.78
VAR(1), Lasso	1	0.37	0.61	0.37	3.34	3.32	1.72	11.98	5.93	10.21
	5	0.44	0.73	0.43	3.77	3.64	2.25	14.25	6.82	12.27
	20	0.69	0.96	0.68	4.37	4.03	3.16	19.98	8.11	18.07
VAR(1), Lasso	1	0.34	0.55	0.33	3.08	3.04	1.76	11.26	5.4	9.72
Post Lasso OLS	5	0.45	0.73	0.44	3.8	3.68	2.23	14.39	6.87	12.36
	20	0.61	0.93	0.6	4.34	4.09	2.94	18.55	8.06	16.43
VAR(1), adaptive Lasso	1	0.37	0.62	0.37	3.46	3.44	1.81	12.21	6.07	10.4
Initial estimator: Lasso	5	0.44	0.74	0.44	3.88	3.78	2.32	14.49	6.93	12.52
	20	0.62	0.98	0.61	4.45	4.18	3.13	19.44	8.38	17.3
VAR(1), Lasso	1	0.35	0.58	0.35	3.25	3.25	1.42	11.31	5.76	9.48
Log-matrix transform	5	0.42	0.73	0.41	3.58	3.58	1.62	13.26	6.65	11.2
	20	0.48	0.94	0.47	4.02	4.02	1.81	15.27	8.04	12.64
VAR(1), Lasso	1	0.37	0.61	0.37	3.34	3.32	1.77	12.44	5.93	10.22
Including S&P 500	5	0.44	0.74	0.43	3.79	3.65	2.33	14.77	6.84	12.31
	20	0.68	0.95	0.67	4.37	4.02	3.16	20.46	8.08	17.96

Forecasting Results

		Α	MedAl	FΕ	A	MaxAI	FΕ		ℓ_2	
Model	h	Α	D	0	Α	D	0	Α	D	
No-Change	1	0.33	0.57	0.33	3.53	3.53	1.47	11.22	5.98	
Censored	5	0.46	0.79	0.45	4.51	4.51	1.91	15.02	7.89	
	20	0.58	0.98	0.57	5.12	5.12	2.22	18.05	9.25	
DCC	1	0.56	0.95	0.55	8.40	8.36	4.28	22.37	12.40	
$\mathrm{EWMA}(\lambda=0.96)$	1	0.88	1.08	0.88	8.07	8.03	4.55	28.89	12.55	
VAR(20), Lasso	1	0.35	0.57	0.35	3.19	3.16	1.62	11.35	5.59	
	5	0.41	0.65	0.4	3.54	3.46	2.01	13.09	6.28	
	20	0.54	0.84	0.53	4.03	3.87	2.56	16.29	7.44	
VAR(20), Lasso	1	0.33	0.52	0.32	3.01	2.92	1.76	10.88	5.09	
Post Lasso OLS	5	0.42	0.66	0.41	3.56	3.48	2.1	13.43	6.31	
	20	0.49	0.79	0.47	4.02	3.9	2.38	15.29	7.27	
VAR(20), adaptive Lasso	1	0.36	0.59	0.35	3.45	3.44	1.61	11.76	5.98	
Initial estimator: Lasso	5	0.43	0.69	0.42	3.75	3.72	2.01	13.62	6.66	
	20	0.58	0.93	0.57	4.16	4.04	2.68	17.49	8.03	
VAR(20), Lasso	1	0.36	0.57	0.35	3.16	3.16	1.39	11.22	5.59	
Log-matrix transform	5	0.4	0.66	0.39	3.42	3.42	1.54	12.53	6.22	
	20	0.46	0.84	0.45	3.81	3.8	1.73	14.37	7.36	
VAR(20), Lasso	1	0.35	0.57	0.35	3.19	3.16	1.64	11.78	5.59	
Including S&P 500	5	0.41	0.65	0.4	3.54	3.46	2.01	13.55	6.28	
	20	0.54	0.84	0.53	4	3.85	2.54	16.83	7.4	

The investor's problem at $t = t_0, ..., T - 1$ is to select a vector of weights for period t + 1 based solely on information up to time t:

$$\widehat{\omega}_{t+1} = \arg\min_{\omega_{t+1}} \omega'_{t+1} \widehat{\Sigma}_{t+1} \omega_{t+1}$$

s.t. $\omega'_{t+1} \widehat{\mu}_{t+1} = \mu_{\text{target}}$
 $\sum_{i=1}^{n} \omega_{it+1} = 1$
 $\sum_{i=1}^{n} |\omega_{it+1}| |\mathbf{I}(\omega_{it} < 0) \le 0.30$
 $|\omega_{it+1}| < 0.20,$

where ω_{t+1} is an $n \times 1$ vector of portfolio weights, μ_{target} is the target expected rate of return from t to t+1 and $I(\cdot)$ is an indicator function.

Model	VAR(1)					V	AR(20)	No-Change	DCC	EWM.	
Estimator:	Lasso	Post Lasso	adaLasso	Lasso	Lasso	Post Lasso	adaLasso	Lasso			
Statistic		OLS	Init: Lasso	(Log Mat)		OLS	Init: Lasso	(Log Mat)	Censored		
Average weight	0.05	0.05	0.05	0.05	0.05	0.05	0.05	0.05	0.05	0.05	0.05
Max weight	0.20	0.20	0.20	0.20	0.20	0.20	0.20	0.20	0.20	0.20	0.20
Min weight	-0.20	-0.20	-0.20	-0.20	-0.20	-0.20	-0.20	-0.20	-0.20	-0.20	-0.20
Average leverage	0.28	0.28	0.28	0.29	0.29	0.29	0.29	0.29	0.29	0.29	0.29
Proportion of leverage	0.22	0.23	0.22	0.27	0.22	0.22	0.22	0.28	0.24	0.24	0.23
Average turnover	0.02	0.03	0.03	0.02	0.02	0.02	0.02	0.01	0.03	0.01	0.01
Average return ($\times 10^{-4}$)	2.58	2.72	3.58	5.89	2.85	2.96	2.64	6.27	1.99	0.40	0.14
Accumulated return	10.07	10.71	15.16	27.77	11.45	11.93	10.34	30.02	7.13	-0.48	-1.60
Standard deviation $(\times 10^2)$	0.97	0.98	0.98	1.00	0.97	0.99	0.97	0.99	0.97	1.00	0.99
Sharpe ratio (×10 ²)	2.66	2.77	3.65	5.87	2.95	3.01	2.70	6.29	2.04	0.40	0.14
Diversification ratio	1.46	1.46	1.47	1.43	1.46	1.46	1.44	1.43	1.48	1.43	1.43
Economic Value $\gamma = 1$											
No-Change (censored)	1.50	1.83	4.08	10.25	2.21	2.45	1.64	11.32	_	_	_
DCC	5.73	6.07	8.41	14.84	6.46	6.72	5.87	15.95	_	_	_
EWMA	6.40	6.74	9.09	15.56	7.13	7.39	6.54	16.68	-	-	-
Economic Value $\gamma = 5$											
No-Change (censored)	1.54	1.77	4.00	9.92	2.27	2.33	1.64	11.06	_	_	_
DCC	6.13	6.37	8.71	14.89	6.90	6.96	6.24	16.08	-	-	-
EWMA	6.68	6.92	9.27	15.48	7.45	7.51	6.79	16.68	-	-	-
Economic Value $\gamma = 10$											
No-Change (censored)	1.58	1.68	3.91	9.50	2.35	2.17	1.63	10.74	_	-	-
DCC	6.64	6.75	9.08	14.95	7.45	7.26	6.69	16.25	_	_	_
EWMA	7.04	7.15	9.49	15.38	7.85	7.66	7.09	16.69	_	_	_

Empirical Example: Forecasting Even Larger Realized Covariance Matrices

Brito, Diego, Marcelo C. Medeiros and Ruy M. Ribeiro (2018). Forecasting Large Realized Covariance Matrices: The Benefits of Factor Models and Shrinkage. Working paper available at SSRN id 3163668.

The Setup Curse of Dimensionality

- ▶ RC matrices are highly persistent over time, which suggests an AR model of large order p (usually p > 20).
- Σ_t : $n \times n$ realized covariance matrix.
- $\boldsymbol{y}_t = \operatorname{vech}(\boldsymbol{\Sigma}_t)$, such that

$$\boldsymbol{y}_t = \boldsymbol{\omega} + \sum_{i=1}^p \boldsymbol{\Phi}_i y_{t-i} + \boldsymbol{\epsilon}_t, \ t = 1, \dots, T,$$

where:

- Φ_i , i = 1, ..., p are the $q \times q$ matrices with q = n(n+1)/2; - ω is a $q \times 1$ vector of intercepts.
- ▶ n(n+1)/2 equations with a total of n(n+1)(p+1)/2 parameters.

The Setup

Factor Structure

• Excess return on any asset $i, r_{i,t}$:

$$r_{i,t}^e = \beta_{i1,t} f_{1,t} + \dots + \beta_{iK,t} f_{K,t} + \varepsilon_{i,t} = \boldsymbol{\beta}'_{i,t} \boldsymbol{f}_t + \varepsilon_{i,t},$$

$$\boldsymbol{r}_t^e = \boldsymbol{B}'_t \boldsymbol{f}_t + \boldsymbol{\varepsilon}_t,$$

where:

- $f_{1,t}, \cdots, f_{K,t}$ are the excess returns of K factors;
- $\beta_{ik,t}$, $k = 1, \ldots, K$, are factor loadings for asset i;
- $\varepsilon_{i,t}$ is the idiosyncratic error term.
- ► Factors are linear combinations of returns: long-short stock portfolios where stocks are sorted on firm characteristics:

$$oldsymbol{f}_t = oldsymbol{W}_t oldsymbol{r}_t^e oldsymbol{W}_t ext{ is known}$$

► Loadings are time-varying and are given as:

$$\boldsymbol{B}_t = (\boldsymbol{\Sigma}_{f,t})^{-1} \boldsymbol{W}_t' \boldsymbol{\Sigma}_t$$

The Setup Covariance Decomposition

• Under the assumption $\mathbb{E}(\boldsymbol{\varepsilon}_t | \boldsymbol{f}_t) = \boldsymbol{0}$, we have

$$\Sigma_t = \operatorname{cov}(\boldsymbol{B}_t'\boldsymbol{f}_t) + \operatorname{cov}(\boldsymbol{\varepsilon}_t) = \boldsymbol{B}_t'\boldsymbol{\Sigma}_{f,t}\boldsymbol{B}_t + \boldsymbol{\Sigma}_{\boldsymbol{\varepsilon},t}.$$

► By linearity:

$$\boldsymbol{\Sigma}_{f,t} = \operatorname{cov}(\boldsymbol{f}_t) = \operatorname{cov}(\boldsymbol{W}_t'\boldsymbol{r}_t) = \boldsymbol{W}_t'\boldsymbol{\Sigma}_t\boldsymbol{W}_t.$$

► Therefore,

$$\widehat{\boldsymbol{\Sigma}}_{t+1|t} = \widehat{\boldsymbol{B}}_{t+1|t}' \widehat{\boldsymbol{\Sigma}}_{f,t+1|t} \widehat{\boldsymbol{B}}_{t+1|t} + \widehat{\boldsymbol{\Sigma}}_{\boldsymbol{\varepsilon},t+1|t}.$$

Forecasting Methodology: $\widehat{\Sigma}_{t+1|t} = \widehat{B}'_{t+1|t} \widehat{\Sigma}_{f,t+1|t} \widehat{B}_{t+1|t} + \widehat{\Sigma}_{\varepsilon,t+1|t}$ Realized Factor Covariance Matrices

• Vector HAR model for
$$\boldsymbol{y}_{f,t} = \mathsf{vech}[\log M(\boldsymbol{\Sigma}_{f,t})]$$
:

$$oldsymbol{y}_{f,t} = oldsymbol{\omega} + oldsymbol{\Phi}_{\mathsf{day}} oldsymbol{y}_{f,t-1}^{\mathsf{day}} + oldsymbol{\Phi}_{\mathsf{week}} oldsymbol{y}_{f,t-1}^{\mathsf{week}} + oldsymbol{\Phi}_{\mathsf{month}} oldsymbol{y}_{f,t-1}^{\mathsf{month}} + oldsymbol{\epsilon}_t,$$

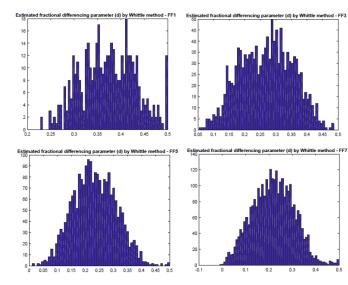
where:

$$- \boldsymbol{y}_{f,t}^{\text{day}} = \operatorname{vech}(\boldsymbol{\Sigma}_{f,t}^{\text{day}}); \, \boldsymbol{y}_{f,t}^{\text{week}} = \operatorname{vech}(\boldsymbol{\Sigma}_{f,t}^{\text{week}}); \, \boldsymbol{y}_{f,t}^{\text{month}} = \operatorname{vech}(\boldsymbol{\Sigma}_{f,t}^{\text{month}}); \\ - \boldsymbol{\Sigma}_{f,t}^{\text{day}} = \log M(\boldsymbol{\Sigma}_{f,t}); \\ - \boldsymbol{\Sigma}_{f,t}^{\text{week}} = \frac{1}{5} [\log M(\boldsymbol{\Sigma}_{f,t}) + \dots + \log M(\boldsymbol{\Sigma}_{f,t-4})]; \text{ and} \\ - \boldsymbol{\Sigma}_{f,t}^{\text{month}} = \frac{1}{22} [\log M(\boldsymbol{\Sigma}_{f,t}) + \dots + \log M(\boldsymbol{\Sigma}_{f,t-21})].$$

► Estimation via LASSO/adaLASSO

- ▶ Penalty parameter is set with the BIC
- ▶ The inverse LASSO estimates (in absolute value) are used as weights for the adaLASSO

Forecasting Methodology: $\widehat{\Sigma}_{t+1|t} = \widehat{B}'_{t+1|t} \widehat{\Sigma}_{f,t+1|t} \widehat{B}_{t+1|t} + \widehat{\Sigma}_{\varepsilon,t+1|t}$ Loadings



Forecasting Methodology: $\widehat{\Sigma}_{t+1|t} = \widehat{B}'_{t+1|t} \widehat{\Sigma}_{f,t+1|t} \widehat{B}_{t+1|t} + \widehat{\Sigma}_{\varepsilon,t+1|t}$ Loadings

► Loading dynamics modeled as a HAR model:

 $\beta_{k,i,t} = \omega + \phi_{\mathsf{day}} \beta_{k,i,t-1}^{\mathsf{day}} + \phi_{\mathsf{week}} \beta_{k,i,t-1}^{\mathsf{week}} + \phi_{\mathsf{month}} \beta_{k,i,t-1}^{\mathsf{month}} + \epsilon_{k,i,t},$

where $\beta_{k,i,t}$ is the (k, i) element of B_t , i.e., the loading of stock *i* on factor *k* at date *t*.

► Coefficients estimated by OLS.

Forecasting Methodology: $\widehat{\Sigma}_{t+1|t} = \widehat{B}'_{t+1|t} \widehat{\Sigma}_{f,t+1|t} \widehat{B}_{t+1|t} + \widehat{\Sigma}_{\varepsilon,t+1|t}$ Residual Covariance

- ► Forecasting ∑_{ε,t} is still subject to the curse of dimensionality
- We assume that $\Sigma_{\varepsilon,t}$ is **block-diagonal** where blocks are defined by industry classification.
- Furthermore, we assume that the dynamics of each block depends **only** on the elements of the same block at t 1
- ► Finally, past covariances are not used as regressors (Callot, Kock, and Medeiros, 2017)

Forecasting Methodology: $\widehat{\Sigma}_{t+1|t} = \widehat{B}'_{t+1|t} \widehat{\Sigma}_{f,t+1|t} \widehat{B}_{t+1|t} + \widehat{\Sigma}_{\varepsilon,t+1|t}$ Residual Covariance

 \blacktriangleright S sectors:

$$oldsymbol{\Sigma}_{oldsymbol{arepsilon},t} = egin{pmatrix} oldsymbol{\Sigma}_{oldsymbol{arepsilon},t} & & & \ & & \ddots & \ & & & oldsymbol{\Sigma}_{oldsymbol{arepsilon},t} \end{pmatrix}$$

•

▶ The dynamics for $y_{\varepsilon,t}^s = \operatorname{vech}[\log M(\Sigma_{\varepsilon,t}^s)], s \in \{1, 2, \dots, S\}$:

$$ig| oldsymbol{y}_{arepsilon,t}^s = oldsymbol{\omega}_{\epsilon}^s + oldsymbol{\Phi}^s oldsymbol{\Lambda}_{arepsilon,t-1}^s + oldsymbol{u}_{arepsilon,t-1}^s,$$

where $\Lambda_{\varepsilon,t-1}^s = \text{diag}[\log M(\Sigma_{\varepsilon,t-1}^s)].$

► LASSO/adaLASSO estimation equation by equation.

Data

Realized Covariance Matrices

- ► The data consists of daily realized covariance matrices of returns for constituents of the S&P 500 index
- ▶ We consider companies that remained in the index and had balance sheet data for the full sample period, totaling 430 stocks
- ► These matrices were constructed from 5-minute returns by composite realized kernel (Lunde et al, 2016 JBES)
- ► Sample period: January 2006 December 2011 (1495 days). Estimation windows with 1,000 observations.
- ▶ Data cleaning: merges and splits.

Data Factors and Sector Classification

- ▶ 6 factors + market are considered: Size (SMB), Value (HML), Gross Profitability, Investment, Asset Growth and Accruals (CRSP/Compustat database)
- ▶ 4 different combinations: 1F(Market), 3F(1F + Size and Value), 5F(3F + Gross Profitability and Investment), and 7F(5F + Asset Growth and Accruals)
- ► Standard Industrial Classification (SIC): 10 sectors

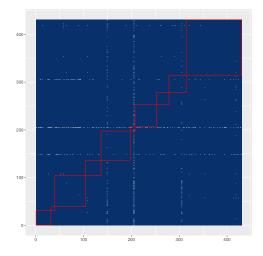
Data Number of Stocks per Sector

Sector	Number of Stocks
Consumer Non-Durables	31
Consumer Durables	8
Manufacturing	65
Oil, Gas, and Coal Extraction	32
Business Equipment	61
Telecommunications	10
Wholesale and Retail	45
Health Care, Medical Equipments, and Drugs	26
Utilities	36
Others	116

Results

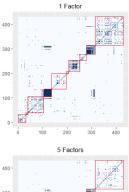
Covariance Structure

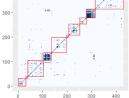
- ▶ The blue dots represent the correlations larger than 0.15 in absolute value in at least 1/3 of the sample days.
- ▶ Red squares represent the groups defined by SIC.

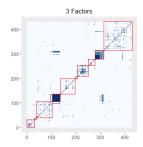


Results

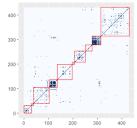
Factor Decomposition and Residual Covariance







7 Factors



Forecasting Results

Forecast Precision for Factor Covariance Matrices

▶ ℓ_2 represents the average ℓ_2 -forecast error over the 473 out-of-sample days, that is,

average
$$\ell_2$$
-forecast error $= \frac{1}{T_2 - T_1 + 1} \sum_{T=T_1}^{T=T_2} ||\widehat{\epsilon}_{T+1}||.$

▶ $\ell_2/\ell_{2,RW}$ represents the ratio of the ℓ_2 -forecast error for other methods to the random walk value.

	ℓ_2	$\ell_2 \ / \ \ell_{2,RW}$					
Model	Random Walk	FHAR	FHAR, Log-matrix				
1F	0.40	0.96 (0.96)	$0.92 \ (0.92)$				
3F	0.44	0.98(0.97)	$0.90 \ (0.90)$				
$5\mathrm{F}$	0.51	$0.95 \ (0.95)$	$0.89 \ (0.89)$				
$7\mathrm{F}$	0.62	0.99(1.04)	$0.86 \ (0.87)$				

Forecasting Results

Forecast Precision for Complete Covariance Matrices

Model (Benchmarks)	$\ell_2/\ell_{2,RW}$	VHAR (Log-matrix)	$\ell_2/\ell_{2,RW}$
RW	1.00	1F, LASSO	0.86
EWMA (Returns)	6.93	3F, LASSO	0.85
BEKK-NL	1.71	5F, LASSO	0.85
DCC-NL	1.71	7F, LASSO	0.85
Block 1F	0.97	1F, adaLASSO	0.86
Block 3F	0.97	3F, adaLASSO	0.85
Block 5F	0.97	5F, adaLASSO	0.85
Block 7F	0.97	7F, adaLASSO	0.85
Random Walk (RW) $\ell_{2,RW}$	341.57		

Statistics for Daily Portfolios - Global Minimum Variance

- Consider the problem of an investor at time $t = t_0, \ldots, T 1$ who wishes to construct a minimum variance portfolio to be held in time t + 1.
- The optimization problem consists of choosing a vector of weights \widehat{w}_{t+1} :

$$\widehat{w}_{t+1} = \operatorname*{arg\,min}_{w_{t+1}} \quad w'_{t+1} \widehat{\Sigma}_{t+1} w_{t+1}$$

subject to $\quad w'_{t+1} \mathbf{1} = 1.$

Statistics for Daily Portfolios - Global Minimum Variance

	RW	Block 1F	Block 3F	Block 5F	Block 7F	EWMA	BEKK-NL	DCC - NL	
Standard Deviation (%)	12.07	8.21	8.29	8.25	8.25	14.62	9.41	10.65	
Lower Partial SD (%)	12.82	8.79	8.94	8.73	8.83	14.90	9.63	11.31	
Avg. Gross Leverage	5.94	3.08	3.14	3.14	3.19	12.55	5.09	4.11	
Prop. of Leverage (%)	44.30	44.40	44.22	44.10	44.11	49.17	45.11	51.73	
Avg. Turnover (%)	1.80	0.75	0.78	0.78	0.80	0.27	0.11	0.21	
Avg. Excess Return (%)	14.20	12.72	14.46	15.37	14.95	3.42	17.98	17.46	
Cumulative Return (%)	29.04	26.42	30.59	32.86	31.82	4.74	39.27	37.58	
Sharpe Ratio	1.18	1.55	1.74	1.86	1.81	0.23	1.91	1.64	
	1 F	actor	3 Fa	3 Factors		5 Factors		ctors	
	VI	IAR	VHAR		VHAR		VHAR		
	(Log :	matrix)	(Log n	natrix)	(Log matrix)		(Log n	natrix)	
	LASSO	aLASSO	LASSO	aLASSO	LASSO	aLASSO	LASSO	aLASSO	
Standard Deviation (%)	8.46	8.42	8.37	8.32	8.29	8.25	8.12	8.09	
Lower Partial SD (%)	8.86	8.81	8.78	8.68	8.57	8.53	8.52	8.51	
Avg. Gross Leverage	2.66	2.67	2.80	2.80	2.82	2.82	2.93	2.93	
Prop. of Leverage (%)	45.89	46.01	44.88	45.03	44.89	45.12	45.26	45.50	
Avg. Turnover (%)	0.20	0.22	0.20	0.22	0.19	0.21	0.20	0.22	
Avg. Excess Return (%)	15.24	15.18	17.69	17.45	18.93	18.61	18.09	17.85	
Cumulative Return (%)	32.49	32.35	38.74	38.13	42.01	41.19	39.85	39.21	
Sharpe Ratio	1.80	1.80	2.11	2.10	2.28	2.26	2.23	2.21	

Statistics for Daily Portfolios - Restricted Minimum Variance

- ▶ Maximum leverage equal to 30% (in some sense, consistent with a 130-30 fund concept in the mutual fund industry).
- Maximum weights on individual stocks: 20% (in absolute value).
- ▶ The problem for an investor at time $t = t_0, ..., T 1$ is then given by

$$\widehat{\boldsymbol{w}}_{t+1} = \underset{\boldsymbol{w}_{t+1}}{\operatorname{arg min}} \quad \boldsymbol{w}'_{t+1} \widehat{\boldsymbol{\Sigma}}_{t+1} w_{t+1}$$
subject to $\boldsymbol{w}'_{t+1} \mathbf{1} = 1$,
$$\sum_{i=1}^{N} |w_{it+1}| I(w_{it} < 0) \le 0.30 \quad \text{and} \quad |w_{it+1}| \le 0.20.$$

Statistics for Daily Portfolios - Restricted Minimum Variance

	RW	Block 1F	Block 3F	Block 5F	Block 7F	EWMA	BEKK-NL	DCC - NL	
Standard Deviation (%)	13.29	13.34	13.20	13.17	13.25	15.28	15.49	14.72	
Lower Partial SD (%)	14.13	13.91	13.66	13.35	13.68	16.47	16.24	15.28	
Avg. Gross Leverage	1.60	1.60	1.60	1.60	1.60	1.60	1.60	1.60	
Prop. of Leverage (%)	1.91	3.11	3.08	3.06	2.93	0.71	0.85	1.41	
Avg. Turnover (%)	0.43	0.40	0.42	0.41	0.42	0.09	0.10	0.11	
Avg. Excess Return (%)	16.72	18.23	19.01	22.42	21.22	13.68	14.24	16.91	
Cumulative Return (%)	34.88	38.74	40.83	50.14	46.79	26.74	27.99	34.86	
Sharpe Ratio	1.26	1.37	1.44	1.70	1.60	0.90	0.92	1.15	
	1 F	actor	3 Fa	ctors	5 Factors		7 Factors		
	VI	HAR	VH	VHAR		VHAR		AR	
	(Log	matrix)	(Log n	natrix)	(Log matrix)		(Log n	natrix)	
	LASSO	aLASSO	LASSO	aLASSO	LASSO	aLASSO	LASSO	aLASSO	
Standard Deviation (%)	13.20	13.37	12.81	12.86	12.57	12.83	12.63	12.75	
Lower Partial SD (%)	13.29	13.64	12.60	12.54	12.54	12.75	12.52	12.62	
Avg. Gross Leverage	1.60	1.60	1.60	1.60	1.60	1.60	1.60	1.60	
Prop. of Leverage (%)	2.46	2.44	2.37	2.38	2.43	2.41	2.27	2.25	
Avg. Turnover (%)	0.22	0.23	0.24	0.24	0.23	0.24	0.22	0.23	
Avg. Excess Return (%)	16.07	19.89	19.72	21.04	20.56	18.93	20.74	19.19	
Cumulative Return (%)	33.30	43.13	42.88	46.43	45.22	40.76	45.67	41.48	
Sharpe Ratio	1.22	1.49	1.54	1.64	1.64	1.48	1.64	1.51	

Statistics for Daily Portfolios - Restricted Minimum Variance (Long Only)

- ► No short-selling.
- The problem for an investor at time $t = t_0, \ldots, T 1$ is then given by

$$\widehat{\boldsymbol{w}}_{t+1} = \underset{\boldsymbol{w}_{t+1}}{\operatorname{arg min}} \quad \boldsymbol{w}'_{t+1} \widehat{\boldsymbol{\Sigma}}_{t+1} \boldsymbol{w}_{t+1} \\
\text{subject to} \quad \boldsymbol{w}'_{t+1} \mathbf{1} = 1, \\
0 \le w_{it+1} \le 0.20.$$

Statistics for Daily Portfolios - Restricted Minimum Variance (Long Only)

	RW	Block 1F	Block 3F	Block 5F	Block 7F	EWMA	BEKK-NL	DCC - NL
Standard Deviation (%)	17.10	17.06	16.96	16.85	16.88	17.74	17.92	17.78
Lower Partial SD (%)	17.56	17.83	17.63	17.49	17.58	18.94	19.16	19.13
Avg. Gross Leverage	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00
Prop. of Leverage (%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
Avg. Turnover (%)	0.17	0.16	0.16	0.16	0.16	0.03	0.03	0.04
Avg. Excess Return (%)	14.29	15.86	16.18	14.98	15.06	20.22	15.85	16.28
Cumulative Return (%)	27.49	31.30	32.15	29.25	29.44	42.18	30.91	32.04
Sharpe Ratio	0.84	0.93	0.95	0.89	0.89	1.14	0.88	0.92
	1 Factor VHAR			3 Factors VHAR		5 Factors VHAR		ctors AR
	(Log	matrix)	(Log n	(Log matrix)		(Log matrix)		natrix)
	LASSO	aLASSO	LASSO	aLASSO	LASSO	aLASSO	LASSO	aLASSO
Standard Deviation (%)	16.96	16.98	16.55	16.59	16.34	16.47	16.31	16.44
Lower Partial Standard Deviation (%)	17.51	17.64	17.29	17.27	16.88	17.10	16.89	17.03
Prop. of Leverage (%)	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
Avg. Turnover (%)	0.08	0.08	0.07	0.08	0.07	0.08	0.07	0.07
Avg. Excess Return (%)	17.60	17.57	17.62	18.04	18.02	18.17	17.13	17.04
Cumulative Return (%)	35.71	35.63	35.95	37.01	37.06	37.38	34.79	34.50
Sharpe Ratio	1.04	1.03	1.06	1.09	1.10	1.10	1.05	1.04